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STOR 435.001 Lecture 15

**Jointly distributed Random Variables - III**

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### Contents of this lecture

1. Conditional pmf/pdf: definition and simple properties.
2. Functions of **two random variables**. Finding joint distributions.

### Brief aside

1. How many of you have heard of the statistical technique “regression”
2. [https://idc9.github.io/stor390/notes/linear\\_regression/linear\\_regression.html](https://idc9.github.io/stor390/notes/linear_regression/linear_regression.html)
3. **Also see**  
<https://www.coursera.org/learn/machine-learning/lecture/Cf8DF/multivariate-gaussian-distribution>
4. **Then see:** <https://www.coursera.org/learn/machine-learning/lecture/DnNr9/anomaly-detection-using-the-multivariate-gaussian-distribution>
5. We will see connections between the above and the next topic “Conditional pmf and pdfs”.

**Conditional distributions:** discrete case:  $X$  and  $Y$  are discrete with joint p.m.f.  $p(x, y)$ :

**Conditional p.m.f. of  $X$  given  $Y = y$  ( $y$  fixed):**

$$p_{X|Y}(x|y) = P(X = x|Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{p(x, y)}{p_Y(y)}$$

for all  $y$  with  $p_Y(y) > 0$ .

**Note:** If  $X$  and  $Y$  are independent, then  $p_{X|Y}(x|y) = p_X(x)$ .

**Conditional d.f. of  $X$  given  $Y = y$  ( $y$  fixed):**

$$F_{X|Y}(a|y) = P(X \leq a|Y = y) = \sum_{x \leq a} p_{X|Y}(x|y)$$

for all  $y$  with  $p_Y(y) > 0$ .

## Important note

In many cases it is easy to calculate marginal pmf of one random variable (say  $Y$ ) and conditional pmf of the other (say  $p_{X|Y}$ ). Then we can get joint pmf by

$$p_{XY}(x, y) = p_Y(y)p_{X|Y}(x|y)$$

# Jointly distributed random variables

**Example:** Let  $X$  be the number of claims submitted to a life-insurance company in April and let  $Y$  be the corresponding number but for May. Suppose the joint pmf of the two random variables is given by

$$p_{X,Y}(x,y) = \frac{1}{2} \left(\frac{1}{2}\right)^{x-1} e^{-x} (1 - e^{-x})^{y-1}, \quad x = 0, 1, \dots, y = 1, 2, \dots$$

Find the conditional pmf of  $Y$  given that there were 2 claims in April.

**Citation:** Motivated by a problem from

<http://faculty.atu.edu/mfinan/actuarieshall/Pbook.pdf>





**Conditional distributions:** continuous case:  $X$  and  $Y$  are jointly continuous with density  $f(x, y)$ :

**Conditional density of  $X$  given  $Y = y$  ( $y$  fixed):**

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

for all  $y$  with  $f_Y(y) > 0$ .

**Note:** If  $X$  and  $Y$  are independent, then  $f_{X|Y}(x|y) = f_X(x)$ .

**Conditional probabilities given  $Y = y$  ( $y$  fixed):**

$$P(X \in A|Y = y) = \int_A f_{X|Y}(x|y) dx.$$

**Conditional d.f. of  $X$  given  $Y = y$  ( $y$  fixed):**

$$F_{X|Y}(a|y) = P(X \leq a|Y = y).$$

# Jointly distributed random variables

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**Example:** Suppose that the joint density of  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} \frac{e^{-x/y} e^{-y}}{y}, & 0 < x < \infty, 0 < y < \infty, \\ 0, & \text{otherwise.} \end{cases}$$

Find  $f_{X|Y}(x|y)$  and  $P(X > 1|Y = y)$ .



# Jointly distributed random variables

**Bivariate Normal distribution:** Jointly continuous random variables  $X$  and  $Y$  are bivariate normal if their density is

$$f(x, y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left( \left(\frac{x-\mu_X}{\sigma_X}\right)^2 + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2 - 2\rho\frac{(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y} \right)},$$

for  $-\infty < x, y < \infty$ , where  $\sigma_X > 0, \sigma_Y > 0, \rho \in (-1, 1), -\infty < \mu_X, \mu_Y < \infty$ . Find  $f_{X|Y}(x|y)$ .

Notation:

$$\begin{pmatrix} X \\ Y \end{pmatrix} = N \left( \begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix}, \begin{pmatrix} \sigma_X^2 & \rho\sigma_X\sigma_Y \\ \rho\sigma_X\sigma_Y & \sigma_Y^2 \end{pmatrix} \right).$$

Here,  $\mu_X = EX, \mu_Y = EY, \sigma_X^2 = Var(X), \sigma_Y^2 = Var(Y)$ . The parameter  $\rho$  accounts for dependence and will be clarified in the next chapter.

This and generalizations: multivariate normal are one of the foundation pieces of modern statistics. Linear Regression based on this fundamental object!



## Example

Suppose we wanted to plot the pdf of the above function. Let us say  $\mu_X = 0$ ,  $\mu_Y = 0$ ,  $\sigma_X^2 = 1$ ,  $\sigma_Y^2 = 1$  and  $\rho = .5$ . Using Mathematica one can plot the above pdf on the region  $\{(x, y) : -2 \leq x \leq 2, -2 \leq y \leq 2\}$

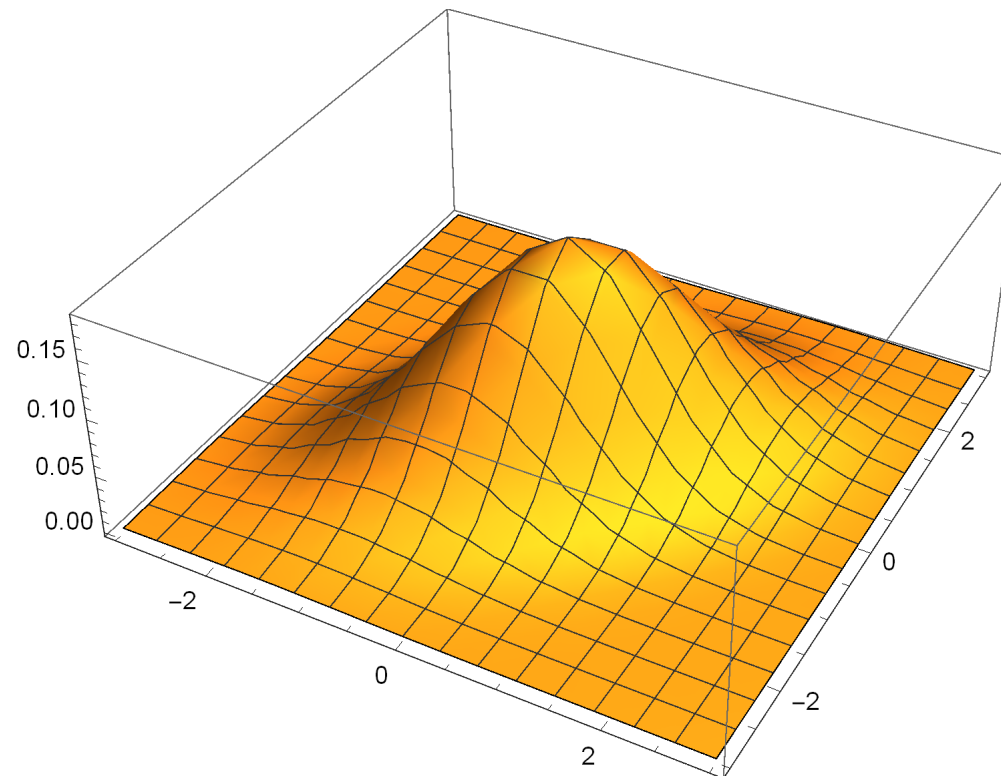
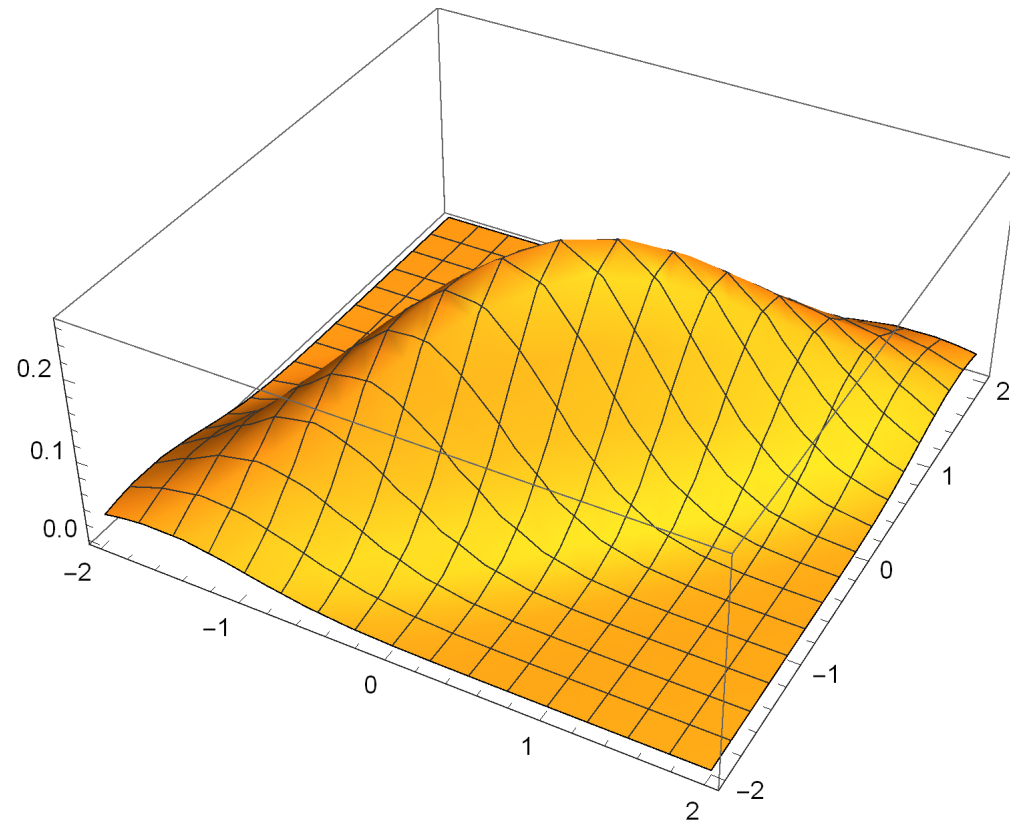


Figure: Using Mathematica



**Figure:** Using Mathematica. Same parameters as above with only change  $\rho = .8$ . Higher “correlation” between  $X$  and  $Y$ . In both cases, the univariate distributions of  $X$  and  $Y$  turn out to be standard normal. In the second case:  $X$  and  $Y$  tend to be “closer” whatever that means. Will see more in the next chapter.

# Jointly distributed random variables

**(Bivariate Normal distribution) contd:**



Look at the exponent in the formula of the Bivariate normal

$$\frac{1}{2(1-\rho^2)} \left( \left( \frac{x-\mu_X}{\sigma_X} \right)^2 + \left( \frac{y-\mu_Y}{\sigma_Y} \right)^2 - 2\rho \frac{(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y} \right)$$

Messy algebra implies that one can write this as

$$= \frac{(x-\mu_X)^2}{2\sigma_X^2} + \frac{1}{2\sigma_Y^2(1-\rho^2)} \left( y-\mu_Y - \rho \frac{\sigma_Y}{\sigma_X} (x-\mu_X) \right)^2$$

Thus

$$f_{XY}(x, y) = \underbrace{\frac{1}{\sigma_X \sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu_X)^2}{\sigma_X^2}}}_{=\text{pdf of}} \cdot \underbrace{\frac{1}{\sigma_Y \sqrt{1-\rho^2} \sqrt{2\pi}} e^{-\frac{1}{2} \frac{\left( y-\mu_Y - \rho \frac{\sigma_Y}{\sigma_X} (x-\mu_X) \right)^2}{\sigma_Y^2 (1-\rho^2)}}}_{=\text{pdf of}}$$

**(Bivariate Normal distribution) contd:**

Thus marginal of  $X$

$$f_X(x) =$$



Conditional pdf of  $Y|X$

$$f_{Y|X}(y|x) = \frac{f_{XY}(x, y)}{f_X(x)} =$$



Thus the distribution of  $Y$  given that  $X = x$  is



## Why is this important?

1. As I said huge parts of statistics are based on extensions of the above to higher dimensions.
2. For our class: Suppose someone came and asked you to simulate

$$\begin{pmatrix} X \\ Y \end{pmatrix} = N \left( \begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix}, \begin{pmatrix} \sigma_X^2 & \rho\sigma_X\sigma_Y \\ \rho\sigma_X\sigma_Y & \sigma_Y^2 \end{pmatrix} \right).$$

3. Step 1: Simulate  $Z_1, Z_2$  standard Normal random variables.
4. Step 2:  $X = \mu_X + \sigma_X Z_1$ .
5. Step 3:

$$Y = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (X - \mu_X) + Z_2 \sigma_Y \sqrt{(1 - \rho^2)}$$

### A related question

How do you simulate standard normal random variables? Assume there are algorithms to simulate  $U(0, 1)$  random variables. This is one motivation for the next topic.

# Joint probability distribution of functions of random variables

Suppose  $X_1, X_2$  are jointly continuous with density  $f(x, y)$ . Consider  $Y_1 = g_1(X_1, X_2)$ ,  $Y_2 = g_2(X_1, X_2)$ , for example,  $g_1(x_1, x_2) = x_1 + x_2$  and  $g_2(x_1, x_2) = x_1 - x_2$ . Suppose the map

$$\begin{pmatrix} y_1 \\ y_2 \end{pmatrix} = \begin{pmatrix} g_1(x_1, x_2) \\ g_2(x_1, x_2) \end{pmatrix}$$

is continuous, differentiable and invertible. What is the joint density of  $Y_1, Y_2$ ?

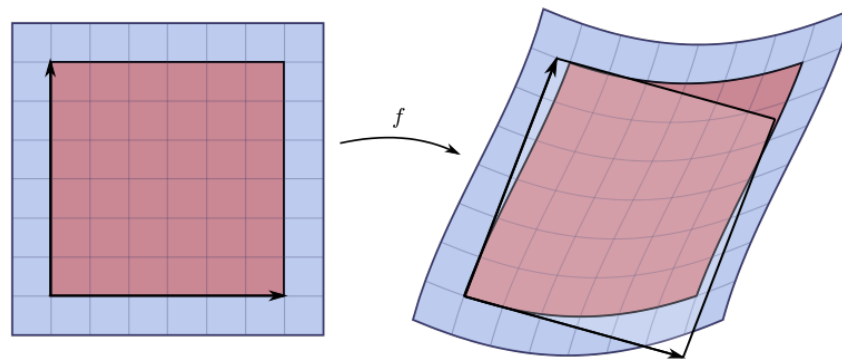


Figure: Jacobian demonstration. Picture from wikipedia.



# Joint probability distribution of functions of random variables contd

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# Jointly distributed random variables (Box Muller transformation)

Let  $(U, T)$  be two independent random variables with  $U \sim \text{Unif}(0, 2\pi)$  and  $T \sim \exp(1)$ . Consider the transformation

$$X = \sqrt{2T} \cos(U), \quad Y = \sqrt{2T} \sin(U)$$

What is the joint distribution of  $(X, Y)$ ?





# Functions of jointly distributed random variables

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**Example:** Let  $X, Y$  be independent uniform  $(0, 1)$  random variables. Find the distribution (pdf) of  $U = XY, V = Y$ .

