# STOR 435.001 Lecture 14 <br> Jointly distributed Random Variables - II 

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## !

Discrete case: Independence is equivalent to

$$
p(x, y)=p_{X}(x) p_{Y}(y), \quad \text { all } x, y
$$

where $p(x, y)$ is the joint p.m.f. of $X$ and $Y, p_{X}$ is the p.m.f. of $X, p_{Y}$ is the p.m.f. of $Y$.

Jointly continuous case: Independence is equivalent to

$$
f(x, y)=f_{X}(x) f_{Y}(y), \quad \text { all } x, y
$$

where $f(x, y)$ is the joint p.d.f. of $X$ and $Y, f_{X}$ is the p.d.f. of $X, f_{Y}$ is the p.d.f. of $Y$.

## Problem from Final Exam 2011

Suppose that 2 balls are chosen without replacement from an urn containing 5 white and 8 red balls. Let $X_{1}$ equal 1 if the 1 st ball selected is red, and let it equal 0 otherwise. Let $X_{2}$ equal 1 if the 2 nd ball selected is red, and let it equal 0 otherwise.

1. Give the joint probability mass function of $X_{1}$ and $X_{2}$.
2. Are $X_{1}$ and $X_{2}$ independent? (Provide a mathematical argument.)

## Jointly distributed random variables

Example: Suppose that the number of people who enter Starbucks on a given day is a Poisson random variable with parameter $\lambda$. Each person who enters the post office orders a single Latte with probability $p$ and orders something else with probability $1-p$. Let $L$ and $N L$ be the total number of single order Latte and Non single order Lattes in a day. Show that these are independent Poisson random variables with respective parameters $\lambda p$ and $\lambda(1-p)$.

Problem: Two points are selected randomly on a line of length $L$ so as to be on opposite sides of the midpoint of the line. (In other words, the two points $X$ and $Y$ are independent random variables such that $X$ is uniformly distributed over $(0, L / 2)$ and $Y$ is uniformly distributed over $(L / 2, L)$.) Find the probability that the distance between the two points is greater than $L / 3$.

Problem continued

Jointly distributed random variables

Example: Let $X=U(0,1)$ and $Y=\operatorname{Exp}(1)$ be independent. Find cdf and pdf of $Z=X+Y$.
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Problem continued

Independence: Random variables $X_{1}, X_{2}, \ldots, X_{n}$ are independent if, for any sets $A_{1}, A_{2}, \ldots, A_{n}$,

$$
\begin{gathered}
P\left(X_{1} \in A_{1}, X_{2} \in A_{2}, \ldots, X_{n} \in A_{n}\right) \\
=P\left(X_{1} \in A_{1}\right) P\left(X_{2} \in A_{2}\right) \ldots P\left(X_{n} \in A_{n}\right)
\end{gathered}
$$

(that is, the events $\left\{X_{1} \in A_{1}\right\},\left\{X_{2} \in A_{2}\right\}, \ldots,\left\{X_{n} \in A_{n}\right\}$ are independent). For example, in the jointly continuous case, this is equivalent to

$$
f\left(x_{1}, x_{2}, \ldots, x_{n}\right)=f_{X_{1}}\left(x_{1}\right) f_{X_{2}}\left(x_{2}\right) \ldots f_{X_{n}}\left(x_{n}\right), \quad \text { all } x_{1}, x_{2}, \ldots, x_{n}
$$

where $f\left(x_{1}, x_{2}, \ldots, x_{n}\right)$ is the joint p.d.f. and $f_{X_{i}}\left(x_{i}\right)$ is a marginal p.d.f. of $X_{i}$.

- Now suppose we had a bunch of independent random variables say $X_{1}, X_{2}, \ldots X_{n}$.
- One important way to combine these is adding them i.e. making a new random variable $S_{n}=X_{1}+X_{2}+\cdots X_{n}$.
- Why important?
- Setting one: We take a random sample of 100 UNC voters. Let $X_{i}$ be a 1 if $i$-th individual in sample is voting for Hillary and $X_{i}=0$ otherwise. So now we have $X_{1}, X_{2}, \ldots, X_{100}$. The total number of people in your sample who are voting for Hillary is?
- Setting two: Researchers are comparing study habits in two universities (call these $U$ and $D$ ) and in particular want to study the average amount of time juniors spend studying per day. They take a sample of 40 students from $U$ and 30 from D and see how much time they spend studying in a week. Let $X_{1}, X_{2}, \ldots X_{40}$ denote the respective amount of time for the students at $U$ and $Y_{1}, Y_{2}, \ldots Y_{30}$ denote the same but for students at D . What would they base the conclusions of their study on?
- Punchline: sums of random variables super important! I will show you two methods for calculting the distribution of such objects.
- Method 1: Direct method
- Method 2: Moment generating function


## 1

Sums of independent random variables: E.g. jointly continuous case: if $X$ and $Y$ are independent, $X$ has density $f_{X}$ and $Y$ has density $f_{Y}$, then the density of $X+Y$ is:
$f_{X+Y}(a)=$


Name: convolution of $f_{X}$ and $f_{Y}$.

Example 3a: If $X$ and $Y$ are independent random variables, both uniformly distributed on $(0,1)$, calculate the probability density of $X+Y$.
-
$R$ code to simulate this
$x<-$ runif(10^5)
$\mathrm{y}<-$ runif(10^5)
$z<-x+y$
hist(z)

Note: For independent discrete random variables we could carry out similar calculations but not integrating but "summing":
Example: If $X=\operatorname{Pois}\left(\lambda_{1}\right)$ and $Y=\operatorname{Pois}\left(\lambda_{2}\right)$ are independent, then
$P(X+Y=n)=$
家

## Method II for finding distributions of sums of independent random variables

- The previous slides we directly calculated the distributions of sums. Turns out there is another powerful tool to do the same. The tool uses two ingredients.
- Ingredient 1: Moment generating functions $M_{X}(t):=E\left(e^{t X}\right)$.
- Amazing math fact: Moment generating functions uniquely characterize the distribution. So suppose we do not know the distribution of a random variable but have somehow managed to compute the MGF and recognize it as the MGF of a known distribution, this means that the random variable has that distribution.

We have computed earlier:

## MGF of some standard distributions

1. $X=\operatorname{Bin}(n, p)$

$$
M_{X}(t)=\left(1-p+p e^{t}\right)^{n}
$$

2. $X=\operatorname{Pois}(\lambda)$

$$
M_{X}(t)=e^{\lambda\left(e^{t}-1\right)} .
$$

3. $X=N\left(\mu, \sigma^{2}\right)$

$$
M_{X}(t)=e^{t \mu+t^{2} \sigma^{2} / 2}
$$

4. $X=$ Gamma with parameters $(\alpha, \lambda)$ then $M_{X}(t)$ is finite only when $t<\lambda$ and then

$$
M_{X}(t)=\left(\frac{\lambda}{\lambda-t}\right)^{\alpha}
$$

## Ingredient 2

Suppose $X$ and $Y$ are independent random variables with marginal pdf $f_{X}$ and $f_{Y}$. Let $g(\cdot), h(\cdot)$ be functions. Show that

$$
\mathbb{E}(g(X) h(Y))=\mathbb{E}(g(X)) \mathbb{E}(h(Y))
$$

The same fact holds for discrete independent random variables and not just 2 but any number $n$ independent random variables.

## Jointly distributed random variables

## Combining the two ingredients

Suppose $X_{1}, X_{2}, \ldots X_{n}$ are independent random variables with the mgf of $X_{i}$ given by $M_{X_{i}}$. Let $S$ be the sum of these random variables namely $S=X_{1}+X_{2}+\cdots X_{n}$. Then MGF of $S$

$$
\begin{aligned}
M_{S}(t) & =E\left(e^{t S}\right)=E\left(e^{t\left(X_{1}+X_{2}+\cdots X_{n}\right)}\right)=E\left(e^{t X_{1}+t X_{2}+\cdots+t X_{n}}\right) \\
& =E\left(e^{t X_{1}} e^{t X_{2}} \cdots e^{t X_{n}}\right) \\
& =E\left(e^{t X_{1}}\right) E\left(e^{t X_{2}}\right) \cdots E\left(e^{t X_{2}}\right) \quad \text { by independence } \\
& =M_{X_{1}}(t) M_{X_{2}}(t) \cdots M_{X_{n}}(t)
\end{aligned}
$$

Punchline: MGF of the sum of a bunch of independent random variables is the product of the individual MGFs.

## Method II of computing distribution of sums of independent rvs

1. Compute the individual MGfs $M_{X_{i}}(t)$.
2. Compute the MGF of the sum $S$ using the above formula.
3. See if $M_{S}$ matches the MGF of one of the known distributions. If it does then $S$ has that distribution.

Jointly distributed random variables

## 4

Let $0<p<1$. If $X_{i}, i=1,2, \ldots, m$ are independent with $X_{i}=\operatorname{Bin}\left(n_{i}, p\right)$ what is the distribution of $\sum_{1}^{m} X_{i}$ ? Note: Same $p$ for all the rvs.

Jointly distributed random variables

## $\triangle$

Let $\lambda_{1}, \lambda_{2}, \ldots \lambda_{m}>0$. If $X_{i}, i=1,2, \ldots, m$ are independent Poisson $\lambda_{i}$ what is the distribution of $\sum_{i=1}^{n} X_{i}$ ?
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If $X_{i}, i=1,2, \ldots, n$ are independent and identically distributed exponential random variables with parameter $\lambda$, what is the distribution of $\sum_{j=1}^{n} X_{i}$ ?


Jointly distributed random variables

## 

If $X_{i}$ are independent normal random variables with parameters $\mu_{i}, \sigma_{i}^{2}$, what is the distribution of $\sum_{i=1}^{n} X_{i}$ ?
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Problem: Motivated by

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https://www.tkiryl.com/Elementarytatistics/Chapter_9.pdf.
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The following is taken almost verbatim from the above site. One of the larger species of tarantulas is the Grammostola mollicoma, whose common name is the Brazilian giant tawny red. A tarantula has two body parts. The anterior part of the body is covered above by a shell, or carapace. From a recent article by F. Costa and F. Perez-Miles titled Reproductive Biology of Uruguayan Theraphosids (The Journal of Arachnology, Vol. 30, No. 3, pp. 571-587):
The carapace length of the adult male G. mollicoma is normally distributed with mean 18.14 mm and standard deviation 1.76 mm .

Suppose you find three such males and measure them. Assume the lengths are independent random variables. Find the probability that the average length of the carpace is greater than 19 mm .


Solution continued.

- We thought about operations on independent random variables.
- Focussed on sums of independent random variables.
- Found that in a number of cases the sum of such random variables can be explicitly evaluated.
- Two methods. Method 1: Direct calculation of the pmf or pdf of the sum.
- Method 2: Using moment generating functions.

